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Sign-consistent estimation in a sparse Poisson model

M. Gomtsyan, C. Lévy-Leduc, S. Ouadah, L. Sansonnet

^aUniversité Paris-Saclay, AgroParisTech, INRAE, UMR MIA Paris-Saclay, 22 place de l'agronomie, Palaiseau, 91120, France

Abstract

In this work, we consider an estimation method in sparse Poisson models inspired by [1] and provide novel sign consistency results under mild conditions.

Keywords: Lasso, sign-consistency, variable selection, Poisson model

1. Introduction

Discrete-valued data arise in diverse applied scientific areas, ranging from finance to molecular biology and epidemiology. For example, as discussed in [2], in molecular biology, non-coding genes are emerging as potential key regulators of the expression of protein-coding genes. Yet, among numerous non-coding genes, only a few are likely to be involved for explaining the expression of the coding genes. Consequently, variable selection will help to identify the relevant non-coding genes by obtaining sparse estimators, meaning that most of them are zero. A popular approach in statistics for performing variable selection is the Lasso proposed by [3]. Besides, [4] showed that Lasso has theoretical guarantees, under some mild conditions. More particularly, Lasso is model selection consistent, meaning that Lasso chooses the true model. However, the consistency results are established in a Gaussian setting, which may not hold for discrete-valued data.

In this work we consider the following sparse Poisson model. Let Y_1, \dots, Y_n be independent random variables such that for all i ,

$$Y_i \sim \text{Poisson}(\lambda_i^*) \quad \text{with} \quad \lambda_i^* = \exp(x_i \boldsymbol{\beta}^*), \quad (1)$$

where x_i is the i th row of a $n \times p$ design matrix \mathbf{X} and $\boldsymbol{\beta}^*$ is a sparse vector of regression coefficients in \mathbb{R}^p . The non-null coefficients correspond to the predictors that are relevant to explain the response. In the following we will consider an estimation method inspired by [1] and provide a novel sign consistency result.

The paper is organised as follows. Firstly, in Section 2 we present the statistical approach for estimating $\boldsymbol{\beta}^*$. Next, in Section 3 we establish its sign-consistency. The detailed proof is available in the Appendix.

2. Statistical approach

To estimate $\boldsymbol{\beta}^*$ defined in Model (1), we shall use the approximation proposed by [1] which consists in maximizing with respect to $\boldsymbol{\beta}$ the second order Taylor approximation of the log-

likelihood l at the current estimate $\tilde{\boldsymbol{\beta}}$, namely:

$$l(\tilde{\boldsymbol{\beta}}) + \sum_{i=1}^n \sum_{k=1}^p x_{ik}(Y_i - \tilde{\lambda}_i)(\beta_k - \tilde{\beta}_k) - \frac{1}{2} \sum_{i=1}^n \sum_{1 \leq k, \ell \leq p} \tilde{\lambda}_i x_{ik}(\beta_k - \tilde{\beta}_k) x_{i\ell}(\beta_\ell - \tilde{\beta}_\ell),$$

where $\tilde{\lambda}_i = \exp(x_i \tilde{\boldsymbol{\beta}})$, which means maximizing

$$\sum_{i=1}^n (Y_i - \tilde{\lambda}_i) x_i (\boldsymbol{\beta} - \tilde{\boldsymbol{\beta}}) - \frac{1}{2} \sum_{i=1}^n \tilde{\lambda}_i (x_i (\boldsymbol{\beta} - \tilde{\boldsymbol{\beta}}))^2 = \sum_{i=1}^n \frac{(Y_i - \tilde{\lambda}_i)}{\sqrt{\tilde{\lambda}_i}} \left(\sqrt{\tilde{\lambda}_i} x_i (\boldsymbol{\beta} - \tilde{\boldsymbol{\beta}}) \right) - \frac{1}{2} \sum_{i=1}^n \left(\sqrt{\tilde{\lambda}_i} x_i (\boldsymbol{\beta} - \tilde{\boldsymbol{\beta}}) \right)^2.$$

This boils down to minimizing

$$\sum_{i=1}^n \left(\sqrt{\tilde{\lambda}_i} x_i (\boldsymbol{\beta} - \tilde{\boldsymbol{\beta}}) - \frac{Y_i - \tilde{\lambda}_i}{\sqrt{\tilde{\lambda}_i}} \right)^2.$$

Minimizing this criterion can be viewed as the minimization with respect to $\boldsymbol{\beta}$ of the following least-squares criterion: $\|\mathcal{Y} - \mathcal{X}\boldsymbol{\beta}\|_2^2$ where $\|u\|_2^2 = \sum_{i=1}^n u_i^2$ for a vector $u = (u_1, \dots, u_n)$ in \mathbb{R}^n ,

$$\mathcal{Y} = \mathcal{X}\tilde{\boldsymbol{\beta}} + \tilde{\boldsymbol{\Lambda}}^{-1/2}(\mathbf{Y} - \tilde{\boldsymbol{\lambda}}) \quad \text{and} \quad \mathcal{X} = \tilde{\boldsymbol{\Lambda}}^{1/2}\mathbf{X}, \quad (2)$$

$\tilde{\boldsymbol{\Lambda}}$ denoting the diagonal matrix having the $\tilde{\lambda}_i$'s as diagonal elements, $\tilde{\boldsymbol{\lambda}}$ being a column vector having the $\tilde{\lambda}_i$'s as components and \mathbf{Y} denoting a column vector having the Y_i 's as components.

Thus, in order to obtain a sparse estimation of $\boldsymbol{\beta}^*$, we will focus on finding $\hat{\boldsymbol{\beta}}(\alpha)$ defined for $\alpha > 0$ by:

$$\hat{\boldsymbol{\beta}}(\alpha) = \arg \min_{\boldsymbol{\beta} \in \mathbb{R}^p} \left\{ \|\mathcal{Y} - \mathcal{X}\boldsymbol{\beta}\|_2^2 + \alpha \|\boldsymbol{\beta}\|_1 \right\}, \quad (3)$$

where $\|v\|_1 = \sum_{k=1}^p |v_k|$ for a vector $v = (v_1, \dots, v_p)$ in \mathbb{R}^p .

We shall establish the sign consistency of $\hat{\boldsymbol{\beta}}$ in Theorem (3.1) of the following section.

3. Sign consistency

Let

$$\mathbf{C} = \frac{\mathcal{X}^T \mathcal{X}}{n} \quad \text{and} \quad \mathbf{W} = \frac{\mathcal{X}^T \tilde{\boldsymbol{\epsilon}}}{n}, \quad (4)$$

where A^T denotes the transpose of the matrix A ,

$$\tilde{\boldsymbol{\epsilon}} = (\tilde{\epsilon}_1, \dots, \tilde{\epsilon}_n)^T \quad \text{with} \quad \tilde{\epsilon}_k = \frac{Y_k - \tilde{\lambda}_k}{\sqrt{\tilde{\lambda}_k}} \quad \text{for all} \quad 1 \leq k \leq n. \quad (5)$$

Without loss of generality, suppose that $\boldsymbol{\beta}^* = (\beta_1^*, \dots, \beta_q^*, \beta_{q+1}^*, \dots, \beta_p^*)^T$, where $\beta_j^* \neq 0$ when $1 \leq j \leq q$ and $\beta_j^* = 0$ when $q+1 \leq j \leq p$ and denote

$$\boldsymbol{\beta}_1^* = (\beta_1^*, \dots, \beta_q^*)^T \quad \text{and} \quad \boldsymbol{\beta}_2^* = (\beta_{q+1}^*, \dots, \beta_p^*)^T. \quad (6)$$

Then,

$$\mathbf{C} = \begin{pmatrix} \mathcal{X}_1^T \mathcal{X}_1 / n & \mathcal{X}_1^T \mathcal{X}_2 / n \\ \mathcal{X}_2^T \mathcal{X}_1 / n & \mathcal{X}_2^T \mathcal{X}_2 / n \end{pmatrix} = \begin{pmatrix} C_{11} & C_{12} \\ C_{21} & C_{22} \end{pmatrix} \quad \text{and} \quad \mathbf{W} = \begin{pmatrix} W_1 \\ W_2 \end{pmatrix}. \quad (7)$$

Theorem 3.1. Assume that Y_1, \dots, Y_n are independent random variables such that for all i , $Y_i \sim \text{Poisson}(\lambda_i^*)$ with $\lambda_i^* = \exp(x_i \boldsymbol{\beta}^*)$, where x_i is the i th row of a design matrix \mathbf{X} and $\boldsymbol{\beta}^*$ is defined in (6). Assume also that there exist positive constants $M_1, M_2, M_3, M_4, M_5, M_6, M_7$ and c_1 such that $0 < c_1 \leq 1$, that p does not depend on n and that the following assumptions hold.

- (T1) For all k in $\{1, \dots, n\}$, $\|x_k\|_2 \leq M_1$ and for all ℓ in $\{1, \dots, p\}$, $\|x^{(\ell)}\|_2 \leq M_7$, where $x^{(\ell)}$ denotes the ℓ th column of \mathbf{X} .
- (T2) $\tilde{\boldsymbol{\beta}}$ is a preliminary estimator of $\boldsymbol{\beta}^*$ such that $|\beta_i^* - \tilde{\beta}_i| = O_p(1/n)$, as n tends to infinity, for all $i = 1, \dots, p$.
- (T3) With a probability tending to 1 as n tends to infinity, $\lambda_{\min}(C_{11}) \geq M_2$, where $\lambda_{\min}(A)$ denotes the smallest eigenvalue of the matrix A and C_{11} is defined in (7).
- (T4) With a probability tending to 1 as n tends to infinity, $\lambda_{\max}(C_{12}) \leq M_3$, $\lambda_{\max}(C_{21}) \leq M_4$, and $\lambda_{\max}(C_{22}) \leq M_5$, where $\lambda_{\max}(A)$ denotes the largest eigenvalue of the matrix A .
- (T5) $n^{\frac{1-c_1}{2}} \min_{1 \leq i \leq p} |\beta_i^*| \geq M_6$.

Let us also suppose that the following condition called strong irrepresentable condition holds: there exists $\tau > 2/3$ such that

$$|C_{21} C_{11}^{-1} \text{sign}(\boldsymbol{\beta}_1^*)| \leq 1 - \tau, \quad (8)$$

with a probability tending to 1 when n tends to infinity, where the inequality has to be understood component by component. Then, for all $\alpha = \alpha_n$ such that $\alpha_n = O\left(n^{\frac{c_2+1}{2}}\right)$, where $0 < c_2 < c_1 \leq 1$, $\hat{\boldsymbol{\beta}}$ defined in (3) satisfies

$$\mathbb{P}\left(\text{sign}(\hat{\boldsymbol{\beta}}) = \text{sign}(\boldsymbol{\beta}^*)\right) \rightarrow 1, \quad \text{when } n \rightarrow \infty.$$

The proof of Theorem 3.1 relies on the following proposition.

Proposition 3.2. Under the assumptions of Theorem 3.1, let

$$R_1 = (C(\boldsymbol{\beta}^* - \tilde{\boldsymbol{\beta}}))_1 \quad \text{and} \quad R_2 = (C(\boldsymbol{\beta}^* - \tilde{\boldsymbol{\beta}}))_2. \quad (9)$$

Then,

$$\mathbb{P}\left(\text{sign}(\hat{\boldsymbol{\beta}}) = \text{sign}(\boldsymbol{\beta}^*)\right) \geq \mathbb{P}(A_n \cap B_n),$$

where

$$A_n = \left\{ |C_{11}^{-1} W_1| < |\boldsymbol{\beta}_1^*| - \frac{\alpha}{2n} |C_{11}^{-1} \text{sign}(\boldsymbol{\beta}_1^*)| - |C_{11}^{-1} R_1| \right\}$$

and

$$B_n = \left\{ |C_{21} C_{11}^{-1} W_1 - W_2| \leq \frac{\alpha}{2n} \left(1 - |C_{21} C_{11}^{-1} \text{sign}(\boldsymbol{\beta}_1^*)|\right) - |C_{21} C_{11}^{-1} R_1 - R_2| \right\}.$$

The proof of Proposition 3.2 is in Appendix A.1.

Proving Theorem 3.1 consists in showing that $\mathbb{P}(A_n^c)$ and $\mathbb{P}(B_n^c)$ go to zero as n tends to infinity where S^c denotes the complementary set of the set S . The proof is given in Appendix A.2.

Appendix A. Proofs

Appendix A.1. Proof of Proposition 3.2

First observe that we get from (2) and (5) that

$$\mathcal{Y} = \mathcal{X}\tilde{\boldsymbol{\beta}} + \tilde{\boldsymbol{\varepsilon}}. \quad (\text{A.1})$$

Let us denote $\hat{\boldsymbol{\beta}}$ the estimator $\hat{\boldsymbol{\beta}}(\alpha)$ defined in (3), it satisfies the following Karush–Kuhn–Tucker conditions described in [5, Section 4.2.2]:

$$\begin{aligned} (\mathcal{X}^T(\mathcal{Y} - \mathcal{X}\hat{\boldsymbol{\beta}}))_i &= \frac{\alpha}{2} \text{sign}(\hat{\beta}_i), & \text{if } \hat{\beta}_i \neq 0, \\ |(\mathcal{X}^T(\mathcal{Y} - \mathcal{X}\hat{\boldsymbol{\beta}}))_i| &\leq \frac{\alpha}{2}, & \text{if } \hat{\beta}_i = 0, \end{aligned}$$

which can be rewritten as follows by using (4) and (A.1)

$$(\mathbf{C}(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}^*) - \mathbf{W} + \mathbf{C}(\boldsymbol{\beta}^* - \tilde{\boldsymbol{\beta}}))_i = -\frac{\alpha}{2n} \text{sign}(\hat{\beta}_i), \quad \text{if } \hat{\beta}_i \neq 0, \quad (\text{A.2})$$

$$|(\mathbf{C}(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}^*) - \mathbf{W} + \mathbf{C}(\boldsymbol{\beta}^* - \tilde{\boldsymbol{\beta}}))_i| \leq \frac{\alpha}{2n}, \quad \text{if } \hat{\beta}_i = 0. \quad (\text{A.3})$$

If A_n holds then

$$-|\boldsymbol{\beta}_1^*| < C_{11}^{-1}W_1 - \frac{\alpha}{2n}C_{11}^{-1}\text{sign}(\boldsymbol{\beta}_1^*) - C_{11}^{-1}R_1 < |\boldsymbol{\beta}_1^*|. \quad (\text{A.4})$$

Let $\check{\boldsymbol{\beta}} = (\check{\boldsymbol{\beta}}_1^T, \mathbf{0}^T)^T$ where

$$\check{\boldsymbol{\beta}}_1 = \boldsymbol{\beta}_1^* + C_{11}^{-1}W_1 - \frac{\alpha}{2n}C_{11}^{-1}\text{sign}(\boldsymbol{\beta}_1^*) - C_{11}^{-1}R_1. \quad (\text{A.5})$$

By (A.4), $|\check{\boldsymbol{\beta}}_1 - \boldsymbol{\beta}_1^*| < |\boldsymbol{\beta}_1^*|$ which implies that $\text{sign}(\check{\boldsymbol{\beta}}_1) = \text{sign}(\boldsymbol{\beta}_1^*)$. Hence, $\check{\boldsymbol{\beta}}$ satisfies (A.2).

If B_n holds then

$$\left| C_{21} \left(C_{11}^{-1}W_1 - C_{11}^{-1}R_1 - \frac{\alpha}{2n}C_{11}^{-1}\text{sign}(\boldsymbol{\beta}_1^*) \right) - W_2 + R_2 \right| \leq \frac{\alpha}{2n},$$

which by (A.5) corresponds to (A.3) for $\check{\boldsymbol{\beta}}$ and concludes the proof of Proposition 3.2.

Appendix A.2. Proof of Theorem 3.1

Let us first prove that $\mathbb{P}(A_n^c)$ tends to zero as n tends to infinity. By denoting

$$\boldsymbol{\xi} = (\xi_1, \dots, \xi_q)^T = C_{11}^{-1}W_1 \quad \text{and} \quad \mathbf{b} = (b_1, \dots, b_q)^T = C_{11}^{-1}\text{sign}(\boldsymbol{\beta}_1^*), \quad (\text{A.6})$$

we get that

$$\begin{aligned} \mathbb{P}(A_n^c) &= \mathbb{P} \left(|C_{11}^{-1}W_1| + |C_{11}^{-1}R_1| + \frac{\alpha}{2n} |C_{11}^{-1}\text{sign}(\boldsymbol{\beta}_1^*)| \geq |\boldsymbol{\beta}_1^*| \right) \\ &\leq \sum_{j=1}^q \mathbb{P} \left(|\xi_j| + |(C_{11}^{-1}R_1)_j| + \frac{\alpha}{2n}|b_j| \geq |\beta_j^*| \right) \\ &\leq \sum_{j=1}^q \left\{ \mathbb{P} \left(|\xi_j| \geq \frac{|\beta_j^*|}{3} \right) + \mathbb{P} \left(|(C_{11}^{-1}R_1)_j| \geq \frac{|\beta_j^*|}{3} \right) + \mathbb{P} \left(\frac{\alpha}{2n}|b_j| \geq \frac{|\beta_j^*|}{3} \right) \right\}. \end{aligned}$$

By the Cauchy-Schwarz inequality, we get that for all j in $\{1, \dots, q\}$

$$|b_j| \leq \sum_{j=1}^q |b_j| \leq \sqrt{q} \|b\|_2 = \sqrt{q} \|C_{11}^{-1} \text{sign}(\boldsymbol{\beta}_1^*)\|_2 \leq q \|C_{11}^{-1}\|_2 = q \lambda_{\max}(C_{11}^{-1}).$$

By using (T5), (T3), $\alpha = O(n^{(c_2+1)/2})$ and $0 < c_2 < c_1 \leq 1$, we obtain that

$$\mathbb{P}(A_n^c) \leq \sum_{j=1}^q \mathbb{P}\left(|\xi_j| \geq \frac{M_6 n^{\frac{c_1-1}{2}}}{3}\right) + \sum_{j=1}^q \mathbb{P}\left(|(C_{11}^{-1} R_1)_j| \geq \frac{M_6 n^{\frac{c_1-1}{2}}}{3}\right) + o(1). \quad (\text{A.7})$$

Let us first prove that the second term in the r.h.s of (A.7) tends to 0 as n tends to infinity. Observing that R_1 defined in (9) satisfies: $R_1 = C_{11}(\boldsymbol{\beta}_1^* - \tilde{\boldsymbol{\beta}}_1) + C_{12}(\boldsymbol{\beta}_2^* - \tilde{\boldsymbol{\beta}}_2)$, we get by the Cauchy-Schwarz inequality that

$$|(C_{11}^{-1} R_1)_j| \leq \sqrt{q} \|C_{11}^{-1} R_1\|_2 \leq \sqrt{q} \|\boldsymbol{\beta}_1^* - \tilde{\boldsymbol{\beta}}_1\|_2 + \sqrt{q} \|C_{11}^{-1} C_{12}(\boldsymbol{\beta}_2^* - \tilde{\boldsymbol{\beta}}_2)\|_2.$$

By (T2), (T3), (T4), $|(C_{11}^{-1} R_1)_j| = O_P(n^{-1})$, for all j , where the O_P does not depend on j , which proves that the second term in the r.h.s of (A.7) tends to 0 as n tends to infinity.

Let us now prove that the first term in the r.h.s of (A.7) tends to 0 as n tends to infinity. By (A.6), (4) and (5), denoting $\boldsymbol{\lambda}^*$ the column vector of the λ_i^* 's, ξ can be rewritten as follows

$$\begin{aligned} \xi_j &= (C_{11}^{-1} W_1)_j = \left(C_{11}^{-1} \frac{\mathcal{X}_1^T \tilde{\boldsymbol{\epsilon}}}{n}\right)_j = \left(\frac{1}{n} C_{11}^{-1} \mathcal{X}_1^T \tilde{\boldsymbol{\Lambda}}^{-1/2} (\mathbf{Y} - \tilde{\boldsymbol{\lambda}})\right)_j \\ &= \left(\frac{1}{n} C_{11}^{-1} \mathcal{X}_1^T \tilde{\boldsymbol{\Lambda}}^{-1/2} (\mathbf{Y} - \boldsymbol{\lambda}^*)\right)_j + \left(\frac{1}{n} C_{11}^{-1} \mathcal{X}_1^T \tilde{\boldsymbol{\Lambda}}^{-1/2} (\boldsymbol{\lambda}^* - \tilde{\boldsymbol{\lambda}})\right)_j. \end{aligned} \quad (\text{A.8})$$

For all j in $\{1, \dots, q\}$, all k in $\{1, \dots, n\}$, by (2.3.8) of [6], we have that

$$\begin{aligned} \frac{1}{n} \left(C_{11}^{-1} \mathcal{X}_1^T \tilde{\boldsymbol{\Lambda}}^{-1/2}\right)_{jk} &\leq \frac{1}{\sqrt{n}} \left\| \left(C_{11}^{-1} \frac{\mathcal{X}_1^T}{\sqrt{n}}\right)_{jk} \right\| \sup_{k \in \{1, \dots, n\}} (\tilde{\lambda}_k^{-1/2}) \leq \frac{1}{\sqrt{n}} \left\| C_{11}^{-1} \frac{\mathcal{X}_1^T}{\sqrt{n}} \right\|_2 \sup_{k \in \{1, \dots, n\}} (\tilde{\lambda}_k^{-1/2}) \\ &= \frac{1}{\sqrt{n}} \rho \left(\left(C_{11}^{-1} \frac{\mathcal{X}_1^T}{\sqrt{n}}\right)^T \left(C_{11}^{-1} \frac{\mathcal{X}_1^T}{\sqrt{n}}\right) \right)^{1/2} \sup_{k \in \{1, \dots, n\}} (\tilde{\lambda}_k^{-1/2}), \end{aligned} \quad (\text{A.9})$$

where $\rho(A)$ is the spectral radius of the matrix A . Note that, by Theorem 1.3.22 of [7],

$$\rho \left(\left(C_{11}^{-1} \frac{\mathcal{X}_1^T}{\sqrt{n}}\right)^T \left(C_{11}^{-1} \frac{\mathcal{X}_1^T}{\sqrt{n}}\right) \right) = \rho \left(\left(C_{11}^{-1} \frac{\mathcal{X}_1^T}{\sqrt{n}}\right) \left(C_{11}^{-1} \frac{\mathcal{X}_1^T}{\sqrt{n}}\right)^T \right) = \rho \left(C_{11}^{-1} \frac{\mathcal{X}_1^T \mathcal{X}_1}{n} C_{11}^{-1} \right) = \rho(C_{11}^{-1}). \quad (\text{A.10})$$

From ((T1)) and ((T2)), we get that

$$\tilde{\lambda}_k^{-1/2} = \frac{1}{\sqrt{\exp(x_k \tilde{\boldsymbol{\beta}})}} = \frac{1}{\sqrt{\exp(x_k \boldsymbol{\beta}^*) \exp(x_k (\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta}^*))}} = \lambda_k^{*-1/2} (1 + O_P(n^{-1})),$$

where $\lambda_k^{*-1/2} \leq \exp(-x_k \boldsymbol{\beta}^*/2) \leq \exp(\|x_k\|_2 \|\boldsymbol{\beta}^*\|_2/2) \leq \exp(M_1 \|\boldsymbol{\beta}^*\|_2/2)$, by the Cauchy-Schwarz inequality and (T1). Hence,

$$\sup_{k \in \{1, \dots, n\}} (\tilde{\lambda}_k^{-1/2}) = O_P(1), \text{ as } n \rightarrow \infty. \quad (\text{A.11})$$

By (A.9), (A.10) and (A.11)

$$\frac{1}{n} \left(C_{11}^{-1} X_1^T \tilde{\Lambda}^{-1/2} \right)_{jk} = O_P(n^{-1/2}).$$

Moreover, by (T1) and (T2)

$$\begin{aligned} \lambda_k^* - \tilde{\lambda}_k &= \exp(x_k \boldsymbol{\beta}^*) - \exp(x_k \tilde{\boldsymbol{\beta}}) = \exp(x_k \boldsymbol{\beta}^*) (1 - \exp(x_k (\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta}^*))) \\ &= -\exp(x_k \boldsymbol{\beta}^*) \sum_{\ell \geq 1} \frac{(x_k (\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta}^*))^\ell}{\ell!} = O_P(1/n). \end{aligned} \quad (\text{A.12})$$

Thus, the second term in the r.h.s of (A.8) is $O_P(1/\sqrt{n})$. Note that

$$\begin{aligned} \left(\frac{1}{n} C_{11}^{-1} X_1^T \tilde{\Lambda}^{-1/2} (\mathbf{Y} - \boldsymbol{\lambda}^*) \right)_j &= \left(\frac{1}{n} C_{11}^{-1} X_1^T (\mathbf{Y} - \boldsymbol{\lambda}^*) \right)_j \\ &= \left(\frac{1}{n} (C_{11}^*)^{-1} X_1^T (\mathbf{Y} - \boldsymbol{\lambda}^*) \right)_j + \left(\frac{1}{n} (C_{11}^{-1} - (C_{11}^*)^{-1}) X_1^T (\mathbf{Y} - \boldsymbol{\lambda}^*) \right)_j, \end{aligned} \quad (\text{A.13})$$

where $C^* = X^{*T} X^*/n$ with $X^* = (\Lambda^*)^{1/2} \mathbf{X}$, Λ^* being a diagonal matrix having as diagonal entries the λ_k^* 's. Hence, the second term in the r.h.s of (A.13) is bounded by

$$\|C_{11}^{-1} - (C_{11}^*)^{-1}\|_2 \|X_1^T\|_2 \left(\frac{1}{n} \sum_{k=1}^n |Y_k - \lambda_k^*| \right) = \|C_{11}^{-1} - (C_{11}^*)^{-1}\|_2 O_P(1),$$

by Markov's inequality and (T1). Since, $C_{11}^* = C_{11} + (C_{11}^* - C_{11})$, we have

$$(C_{11}^*)^{-1} = (\text{Id} + C_{11}^{-1}(C_{11}^* - C_{11}))^{-1} C_{11}^{-1}$$

and thus by Corollary 5.6.16 of [7], we get that

$$\|C_{11}^{-1} - (C_{11}^*)^{-1}\|_2 = O_P(\|C_{11}^* - C_{11}\|_2) = O_P(1/n^2) \quad (\text{A.14})$$

since by Equation (2.3.8) of [6], the Cauchy-Schwarz inequality and (A.12)

$$\|C_{11}^* - C_{11}\|_2 \leq \frac{q}{n} \max_{1 \leq \ell \leq q} \|x^{(\ell)}\|_2^2 \sup_{1 \leq k \leq n} |\lambda_k^* - \tilde{\lambda}_k| = O_P(1/n^2). \quad (\text{A.15})$$

Thus, the second term in the r.h.s of (A.13) is $O_P(1/n^2)$. To address the first term in the r.h.s of (A.13), we shall use the following result.

Theorem Appendix A.1. (Bernstein's Inequality, [8, Corollary 2.11]) *Let X_1, \dots, X_n be independent real random variables. Suppose that there exist $v > 0$ and $c > 0$ such that $\sum_{k=1}^n \mathbb{E}[X_k^2] \leq v$, and*

$$\sum_{k=1}^n \mathbb{E}[|X_k|^\ell] \leq \frac{\ell!}{2} v c^{\ell-2}$$

for all integers $\ell \geq 3$. Then, for all $t > 0$,

$$\mathbb{P} \left(\left| \sum_{k=1}^n (X_k - \mathbb{E}[X_k]) \right| \geq t \right) \leq 2 \exp \left[- \frac{t^2}{2(v + ct)} \right].$$

Denoting $G = n^{-1}(C_{11}^*)^{-1}X_1^T(\Lambda^*)^{1/2}$, we get that

$$\left(\frac{1}{n}(C_{11}^*)^{-1}X_1^T(\mathbf{Y} - \lambda^*)\right)_j = \sum_{k=1}^n G_{jk} \frac{(Y_k - \lambda_k^*)}{\sqrt{\lambda_k^*}}.$$

Let us now apply the Bernstein's inequality to $X_k = G_{jk}Y_k / \sqrt{\lambda_k^*}$ then

$$\begin{aligned} \sum_{k=1}^n \mathbb{E}[X_k^2] &= \sum_{k=1}^n G_{jk}^2 (1 + \lambda_k^*) \leq n \sup_{1 \leq k \leq n} (1 + \lambda_k^*) \|G\|_2^2 \\ &\leq n \sup_{1 \leq k \leq n} (1 + \lambda_k^*) \rho(GG^T) \leq \sup_{1 \leq k \leq n} (1 + \lambda_k^*) \rho((C_{11}^*)^{-1}). \end{aligned}$$

By Weyl's inequalities [7, Corollary 4.3.15] and (T3), we get that with a probability tending to 1,

$$\lambda_{\min}(C_{11}^*) \geq \lambda_{\min}(C_{11}) + \lambda_{\min}(C_{11}^* - C_{11}) \geq \lambda_{\min}(C_{11}) - \rho(C_{11}^* - C_{11}) \geq M_2 - \|C_{11}^* - C_{11}\|_2,$$

Denoting $\bar{\lambda} = \max(\sup_{1 \leq k \leq n} \lambda_k^*, 1)$ and M_2' the positive constant such that $\rho((C_{11}^*)^{-1}) \leq 1/M_2'$ for large enough n , which exists by (A.15), we get by (T3) that $\nu = 2\bar{\lambda}/M_2'$. Observe that

$$\sum_{k=1}^n \mathbb{E}[|X_k|^\ell] = \sum_{k=1}^n \mathbb{E} \left[\left| \frac{G_{jk}}{\sqrt{\lambda_k^*}} Y_k \right|^\ell \right] = \sum_{k=1}^n \left(\frac{|G_{jk}|}{\sqrt{\lambda_k^*}} \right)^\ell \sum_{i=1}^{\ell} \lambda_k^{*\ell} \left\{ \begin{matrix} \ell \\ i \end{matrix} \right\}, \quad (\text{A.16})$$

where $\left\{ \begin{matrix} \ell \\ i \end{matrix} \right\}$ denotes the Stirling number of the second kind and the last equality is due to the definition of the ℓ -th moment of a Poisson random variable. Then we have, for all $\ell \geq 3$, by Equation (2.3.8) of [6],

$$\begin{aligned} \sum_{k=1}^n \left(\frac{|G_{jk}|}{\sqrt{\lambda_k^*}} \right)^\ell \sum_{i=1}^{\ell} \lambda_k^{*\ell} \left\{ \begin{matrix} \ell \\ i \end{matrix} \right\} &\leq \sum_{k=1}^n \frac{1}{n^{\ell/2}} \rho((C_{11}^*)^{-1})^{\ell/2} \bar{\lambda}^{\ell/2} \sum_{i=1}^{\ell} \left\{ \begin{matrix} \ell \\ i \end{matrix} \right\} \leq n^{1-\ell/2} M_2'^{-\ell/2} \bar{\lambda}^{\ell/2} \ell! \\ &= n^{1-\ell/2} M_2'^{1-\ell/2} \bar{\lambda}^{\ell/2-1} \frac{\ell!}{2} \nu \leq \frac{\ell!}{2} \nu \left(\frac{\sqrt{\bar{\lambda}}}{\sqrt{nM_2'}} \right)^{\ell-2} = \frac{\ell!}{2} \nu c^{\ell-2}, \end{aligned} \quad (\text{A.17})$$

with $c = \sqrt{\bar{\lambda}/(nM_2')}$ since $\sum_{i=1}^{\ell} \left\{ \begin{matrix} \ell \\ i \end{matrix} \right\} \leq \ell!$.

Hence, for $t = n^{(c_1-1)/2}$,

$$\frac{t^2}{2(\nu + ct)} = \frac{n^{c_1-1}}{2(\nu + n^{-1/2} \sqrt{\bar{\lambda}/M_2'} n^{(c_1-1)/2})} = O(n^{c_1/2}),$$

which gives the expected result.

Let us then prove that $\mathbb{P}(B_n^c)$ tends to zero as n tends to infinity. By denoting

$$\zeta = (\zeta_1, \dots, \zeta_{p-q})^T = C_{21} C_{11}^{-1} W_1 - W_2 \quad \text{and} \quad d = (d_1, \dots, d_{p-q})^T = C_{21} C_{11}^{-1} \text{sign}(\beta_1^*), \quad (\text{A.18})$$

we get that

$$\begin{aligned}
\mathbb{P}(\mathcal{B}_n^c) &= \mathbb{P}\left(|C_{21}C_{11}^{-1}W_1 - W_2| + |C_{21}C_{11}^{-1}R_1 - R_2| + \frac{\alpha}{2n}|C_{21}C_{11}^{-1}\text{sign}(\boldsymbol{\beta}_1^*)| > \frac{\alpha}{2n}\right) \\
&\leq \sum_{j=1}^{p-q} \mathbb{P}\left(|\zeta_j| + |(C_{21}C_{11}^{-1}R_1 - R_2)_j| + \frac{\alpha}{2n}|d_j| \geq \frac{\alpha}{2n}\right) \\
&\leq \sum_{j=1}^{p-q} \left\{ \mathbb{P}\left(|\zeta_j| \geq \frac{\alpha}{6n}\right) + \mathbb{P}\left(|(C_{21}C_{11}^{-1}R_1 - R_2)_j| \geq \frac{\alpha}{6n}\right) + \mathbb{P}\left(|d_j| \geq \frac{1}{3}\right) \right\}. \quad (\text{A.19})
\end{aligned}$$

By the strong irrerepresentable condition (8), we get that $\sum_{j=1}^{p-q} \mathbb{P}\left(|d_j| \geq \frac{1}{3}\right) = o(1)$. Let us now prove that the second term in the r.h.s of (A.19) tends to 0 as n tends to infinity. Observing that R_1 defined in (9) satisfies: $R_1 = C_{11}(\boldsymbol{\beta}_1^* - \tilde{\boldsymbol{\beta}}_1) + C_{12}(\boldsymbol{\beta}_2^* - \tilde{\boldsymbol{\beta}}_2)$, we get by the Cauchy-Schwarz inequality that

$$\begin{aligned}
|(C_{21}C_{11}^{-1}R_1)_j| &\leq \sqrt{p-q} \|C_{21}C_{11}^{-1}R_1\|_2 \\
&\leq \sqrt{p-q} \|C_{21}(\boldsymbol{\beta}_1^* - \tilde{\boldsymbol{\beta}}_1)\|_2 + \sqrt{p-q} \|C_{21}C_{11}^{-1}C_{12}(\boldsymbol{\beta}_2^* - \tilde{\boldsymbol{\beta}}_2)\|_2.
\end{aligned}$$

By (T2), (T3) and (T4), $|(C_{21}C_{11}^{-1}R_1)_j| = O_P(n^{-1})$ for all j . Using similar arguments we get $|(R_2)_j| = O_P(n^{-1})$. Then since $\alpha = O(n^{(c_2+1)/2})$, we get that the second term in the r.h.s of (A.19) tends to 0 as n tends to infinity.

Let us now prove that the first term in the r.h.s of (A.19) tends to 0 as n tends to infinity. By (A.18), (4), (5) and (2), ζ can be rewritten as follows for all $j \in \{1, \dots, p-q\}$

$$\zeta_j = (C_{21}C_{11}^{-1}W_1 - W_2)_j = \left(C_{21}\xi - \frac{X_2^T}{n}\tilde{\boldsymbol{\varepsilon}}\right)_j = \left(C_{21}\xi - \frac{X_2^T}{n}(\mathbf{Y} - \boldsymbol{\lambda}^*)\right)_j, \quad (\text{A.20})$$

where we recall that $\boldsymbol{\lambda}^*$ denotes the column vector of the λ_i^* 's and $\xi = C_{11}^{-1}W_1$.

Let us consider the first term in (A.20) and prove that $\mathbb{P}\left(|(C_{21}\xi)_j| \geq \frac{\alpha}{12n}\right)$ tends to 0 as n tends to infinity. Let us note that the term ξ_j is handled previously in the proof concerning A_n^c . Therefore we use the same arguments and adapt it to the term

$$(C_{21}\xi)_j = \left(\frac{1}{n}C_{21}C_{11}^{-1}X_1^T\tilde{\boldsymbol{\Lambda}}^{-1/2}(\mathbf{Y} - \boldsymbol{\lambda}^*)\right)_j + \left(\frac{1}{n}C_{21}C_{11}^{-1}X_1^T\tilde{\boldsymbol{\Lambda}}^{-1/2}(\boldsymbol{\lambda}^* - \tilde{\boldsymbol{\lambda}})\right)_j. \quad (\text{A.21})$$

The second term of (A.21) is still $O_P(1/\sqrt{n})$ since in (A.10), $\rho(C_{21}C_{11}^{-1}C_{21}^T) = \rho(C_{21}C_{11}^{-1}C_{12})$ is bounded by using (T3) and (T4). The first term of (A.21) can be decomposed in the same way as (A.13):

$$\begin{aligned}
&\left(\frac{1}{n}C_{21}C_{11}^{-1}X_1^T\tilde{\boldsymbol{\Lambda}}^{-1/2}(\mathbf{Y} - \boldsymbol{\lambda}^*)\right)_j \\
&= \left(\frac{1}{n}C_{21}^*(C_{11}^*)^{-1}X_1^T(\mathbf{Y} - \boldsymbol{\lambda}^*)\right)_j + \left(\frac{1}{n}(C_{21}C_{11}^{-1} - C_{21}^*(C_{11}^*)^{-1})X_1^T(\mathbf{Y} - \boldsymbol{\lambda}^*)\right)_j. \quad (\text{A.22})
\end{aligned}$$

The second term of (A.22) is bounded in the same manner as the second term of (A.13) was handled. By observing that

$$\begin{aligned} C_{21}C_{11}^{-1} - C_{21}^*(C_{11}^*)^{-1} &= (C_{21} - C_{21}^*)C_{11}^{-1} + C_{21}^*(C_{11}^{-1} - (C_{11}^*)^{-1}) \\ &= (C_{21} - C_{21}^*)C_{11}^{-1} + (C_{21}^* - C_{21})(C_{11}^{-1} - (C_{11}^*)^{-1}) + C_{21}(C_{11}^{-1} - (C_{11}^*)^{-1}), \end{aligned}$$

and by using (A.14), (T1), (T3), (T4) and the fact that

$$\|C_{21}^* - C_{21}\|_2 = O_P(1/n^2), \quad (\text{A.23})$$

where we used the same arguments as in (A.15) we get that $\|C_{21}C_{11}^{-1} - C_{21}^*(C_{11}^*)^{-1}\|_2 = O_P(1/n^2)$. Thus, the second term of (A.22) is $O_P(1/n^2)$. To handle the first term of (A.22), let us apply the Bernstein's inequality to $X_k = \tilde{G}_{jk}Y_k/\sqrt{\lambda_k^*}$ with $\tilde{G} = n^{-1}C_{21}^*(C_{11}^*)^{-1}X_1^T(\Lambda^*)^{1/2}$. Then

$$\sum_{k=1}^n \mathbb{E}[X_k^2] = \sum_{k=1}^n \tilde{G}_{jk}^2(1 + \lambda_k^*) \leq n \sup_{1 \leq k \leq n} (1 + \lambda_k^*) \rho(\tilde{G}\tilde{G}^T) \leq \sup_{1 \leq k \leq n} (1 + \lambda_k^*) \rho(C_{21}^*(C_{11}^*)^{-1}(C_{21}^*)^T).$$

By Weyl's inequalities [7, Corollary 4.3.15]

$$\begin{aligned} \lambda_{\max}(C_{21}^*(C_{11}^*)^{-1}(C_{21}^*)^T) &\leq \lambda_{\max}((C_{21}^* - C_{21})(C_{11}^*)^{-1} - C_{11}^{-1})(C_{21}^*)^T - C_{21}^T)) \\ &+ \lambda_{\max}((C_{21}^* - C_{21})(C_{11}^*)^{-1} - C_{11}^{-1})C_{21}^T) + \lambda_{\max}((C_{21}^* - C_{21})C_{11}^{-1}((C_{21}^*)^T - C_{21}^T)) \\ &+ \lambda_{\max}((C_{21}^* - C_{21})C_{11}^{-1}C_{21}^T) + \lambda_{\max}(C_{21}((C_{11}^*)^{-1} - C_{11}^{-1})(C_{21}^*)^T - C_{21}^T)) \\ &+ \lambda_{\max}(C_{21}((C_{11}^*)^{-1} - C_{11}^{-1})C_{21}^T) + \lambda_{\max}(C_{21}C_{11}^{-1}((C_{21}^*)^T - C_{21}^T)) + \lambda_{\max}(C_{21}C_{11}^{-1}C_{21}^T). \end{aligned}$$

By (A.14), (A.23), (T3) and (T4), we get that for a large enough n , $\rho(C_{21}^*(C_{11}^*)^{-1}(C_{21}^*)^T) \leq M'_3$, where M'_3 is a positive constant. Hence, $\nu = (2\bar{\lambda})/M'_3$. Using the same bounds as in (A.16) and (A.17), we get that $c = \sqrt{\bar{\lambda}/(nM'_3)}$. Thus, with $t = n^{(c_2-1)/2}$ in the Bernstein's inequality

$$\frac{t^2}{2(\nu + ct)} = \frac{n^{c_2-1}}{2(\nu + n^{-1/2}\sqrt{\bar{\lambda}/M'_3}n^{(c_2-1)/2})} = O(n^{c_2/2}).$$

Consequently, we conclude that for $\alpha = O(n^{(c_2+1)/2})$, $\mathbb{P}\left(|(C_{21}\xi)_j| \geq \frac{\alpha}{12n}\right)$ tends to 0 as n tends to infinity.

Finally, let us consider the second term in (A.20) and prove that $\mathbb{P}\left(\left|\left(\frac{X_1^T}{n}(\mathbf{Y} - \lambda^*)\right)_j\right| \geq \frac{\alpha}{12n}\right)$ tends to 0 as n tends to infinity. Denoting $H = n^{-1}X_2^T(\Lambda^*)^{1/2}$, we observe that

$$\left(\frac{1}{n}X_2^T(\mathbf{Y} - \lambda^*)\right)_j = \sum_{k=1}^n H_{jk} \frac{(Y_k - \lambda_k^*)}{\sqrt{\lambda_k^*}}$$

and we apply the Bernstein's inequality to $X_k = H_{jk}Y_k/\sqrt{\lambda_k^*}$. Then we have

$$\begin{aligned} \sum_{k=1}^n \mathbb{E}[X_k^2] &= \sum_{k=1}^n H_{jk}^2(1 + \lambda_k^*) \leq n \sup_{1 \leq k \leq n} (1 + \lambda_k^*) \|H\|_2^2 \\ &\leq n \sup_{1 \leq k \leq n} (1 + \lambda_k^*) \rho(HH^T) \leq \sup_{1 \leq k \leq n} (1 + \lambda_k^*) \rho(C_{22}^*). \end{aligned}$$

By Weyl's inequalities [7, Corollary 4.3.15]

$$\lambda_{\max}(C_{22}^{\star}) \leq \lambda_{\max}(C_{22}) + \lambda_{\max}(C_{22}^{\star} - C_{22}).$$

By using the same arguments as in (A.15) and (T4), we get that $\rho(C_{22}^{\star}) \leq M'_4$ for large enough n and by denoting $\bar{\lambda} = \max(1, \sup_{1 \leq k \leq n} \lambda_k^{\star})$, we get that $\nu = 2\bar{\lambda}M'_4$. Then, using exactly the same argument as before we get

$$\sum_{k=1}^n \mathbb{E}[|X_k|^{\ell}] \leq \frac{\ell!}{2} \nu c^{\ell-2},$$

with $c = \sqrt{\bar{\lambda}/(nM'_4)}$. Hence, for $t = n^{(c_2-1)/2}$ in the Bernstein's inequality, we obtain the expected result, which concludes the proof.

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